

Type-Based Distributed Estimation over Multiaccess Channels

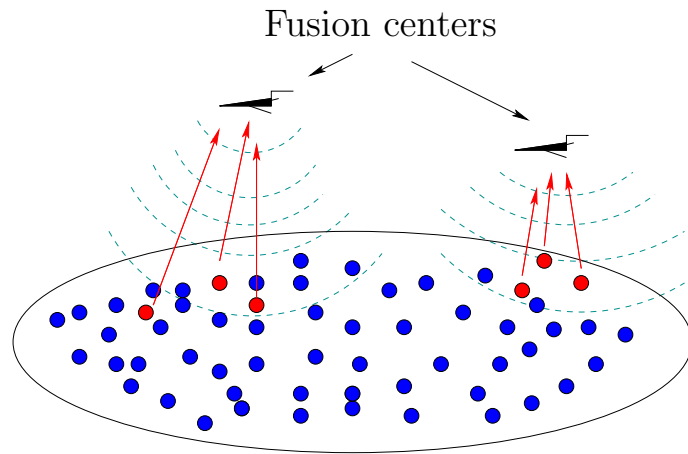
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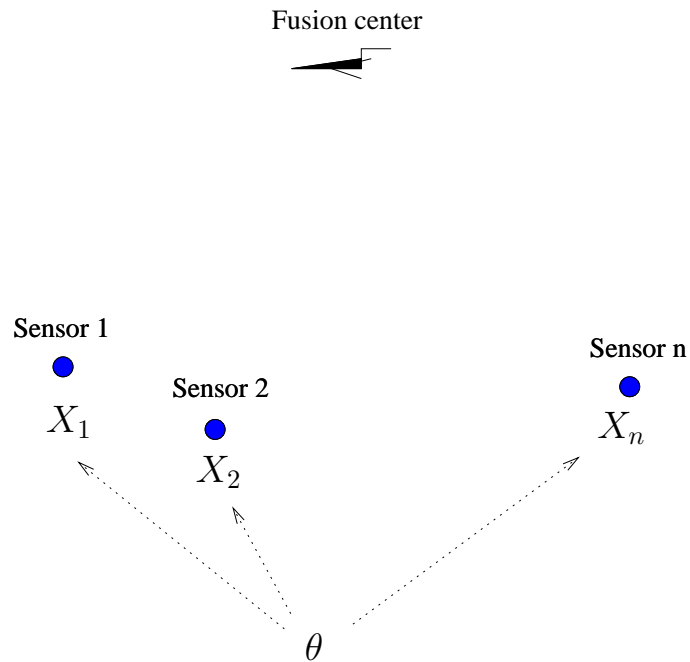
Motivation: Sensor Networks

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- Sense a physical phenomena, transmit it to a fusion center.
- Fusion centers estimate the field parameters, and deliver the estimate.

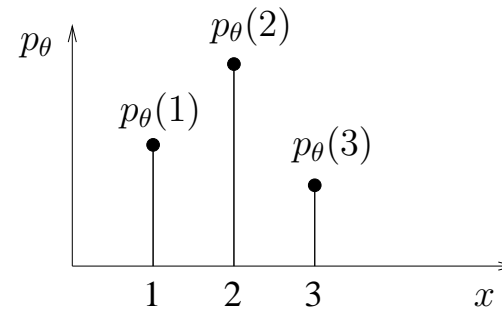
Problem Formulation



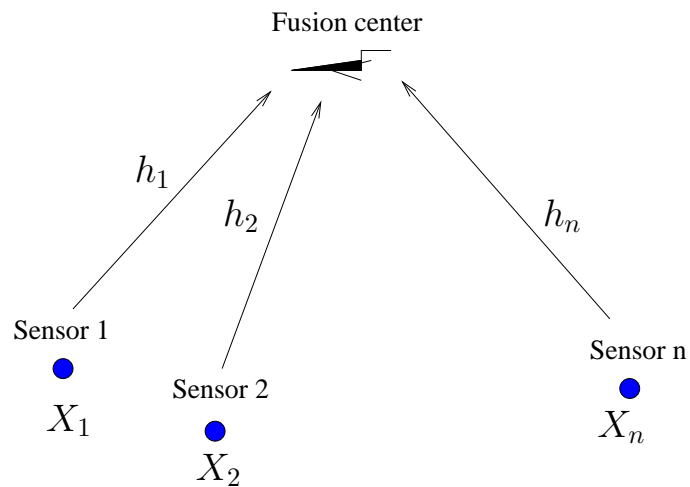
- Estimate $\theta \in \mathbb{R}$.
- Observation X_1, \dots, X_n are i.i.d. conditioned on θ .

Assumptions:

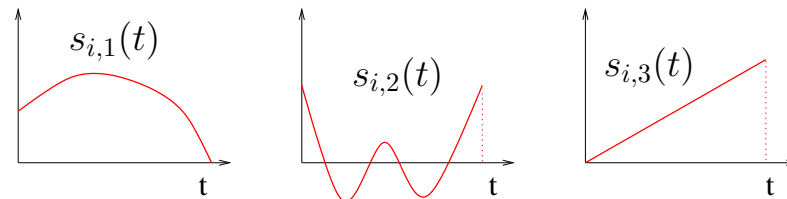
- X_i takes values in $\{1, \dots, k\}$.
- $X_i \sim p_\theta$, where p_θ is a probability mass function.



Problem Formulation Cont'd



- Sensor i has a set of k channel waveforms $s_{i,1}, \dots, s_{i,k}$.
- Upon observing X_i , it transmits s_{i,X_i} .

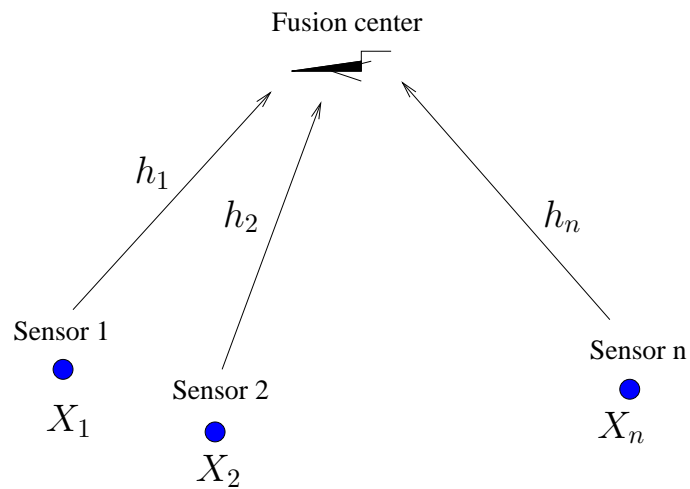


Received signal: $z = \sum_{i=1}^n h_i s_{i,X_i} + w$.

- Channel gains $h_1, \dots, h_n \in \mathbb{R}$ are i.i.d.
- The noise is white $\mathcal{N}(0, \sigma^2)$.
- Energy constraint $\|s_{i,j}\|^2 \leq E$.

Problem Formulation Cont'd

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Estimator:

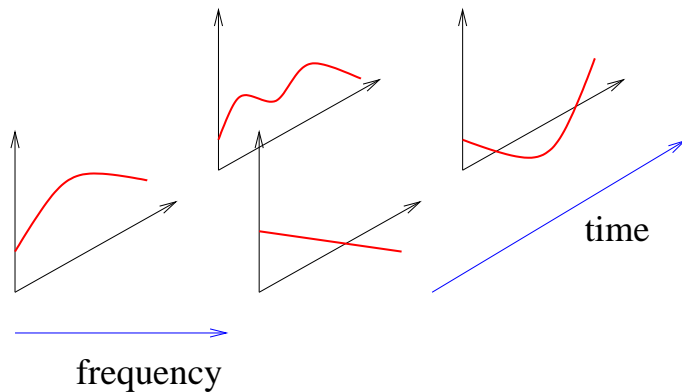
- The parameter θ is estimated based on the received signal z .
- $\hat{\theta}(z)$ is the estimate, $\hat{\theta}$ is the estimator.

Objective:

Design the **channel waveforms** and the **estimator** to minimize the Mean Square Error (MSE) $\mathbb{E}\{(\hat{\theta} - \theta)^2\}$.

Classical Approach

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- Collision among users is “bad!”
- Solution: orthogonalize transmissions.
- Can be done by time/frequency/code division.

Advantages:

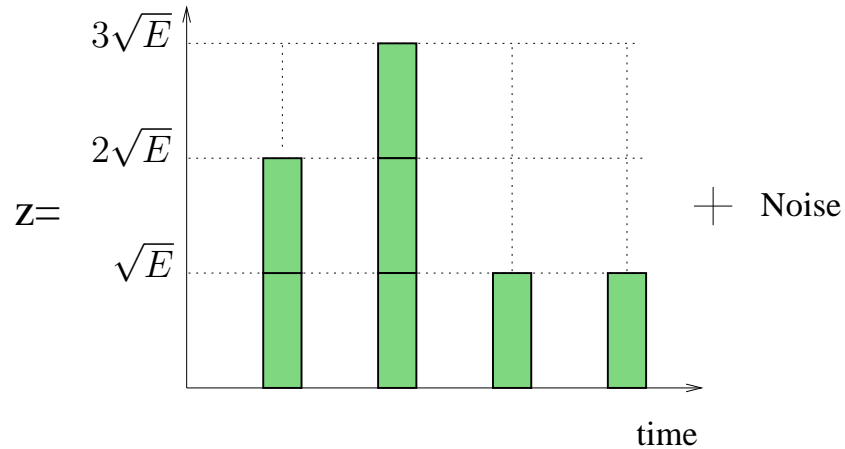
- Allows us to use the standard layered approach.
- Well understood. Rather easy to implement.

Caveat:

- The bandwidth requirement is significant for large n .
- Neglects the dependency among sensor data.

Proposed Approach: TBMA

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- Nodes transmit simultaneously with Pulse Position Modulation (TBMA).
- $s_{i,X_i} = \sqrt{E}\delta_{X_i}$, where $\delta_1, \dots, \delta_k$ are orthonormal pulses.
- When all $h_i = 1$,
 $z = \text{histogram} + \text{noise}$.

Advantages:

- Uses much less bandwidth/time than orthogonal approaches.
- The MSE with TBMA is *asymptotically optimal* as $n \rightarrow \infty$.

Remark:

- Any set of orthonormal $\delta_1, \dots, \delta_k$ can be used.

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- Introduction
 - Performance analysis
 - ▶ **Fundamental limits**
 - TBMA with deterministic h_i
 - TBMA with random h_i
 - Orthogonal allocation
 - Transmitter channel side information
 - Conclusion

Cramer-Rao Bound:

Let $\hat{\theta}$ be an *unbiased* estimator based on X_1, \dots, X_n . Then,

$$\mathbb{E}\{(\hat{\theta} - \theta)^2\} \geq \frac{1}{nI(\theta)}, \quad (1)$$

where $I(\theta) = \sum_{i=1}^k \frac{(dp_{\theta}(i)/d\theta)^2}{p_{\theta}(i)}$ is the *Fisher information* in X_i .

Asymptotic Efficiency:

There exists a class of estimators based on X_1, \dots, X_n satisfying

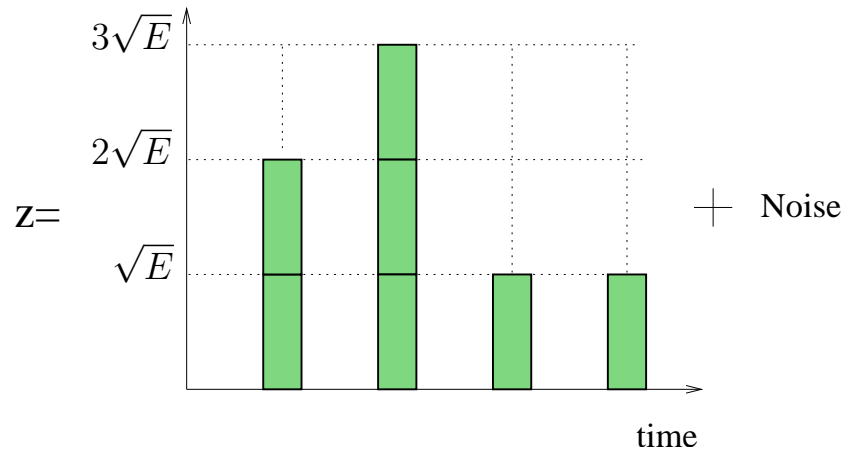
$$\hat{\theta} \doteq \mathcal{N}\left(\theta, \frac{1}{nI(\theta)}\right), \quad \text{for large } n. \quad (2)$$

Notation “ \doteq ” means $\hat{\theta} \xrightarrow{p} \theta$ and $\sqrt{n}(\hat{\theta} - \theta) \xrightarrow{d} \mathcal{N}(0, \frac{1}{I(\theta)})$ as $n \rightarrow \infty$.

Key observations:

- To achieve asymptotic efficiency, an estimator need *not* have access to all data X_1, \dots, X_n .
- Knowledge of a *sufficient statistic* is actually enough.

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- A sufficient statistic is **empirical measure**:

$$\tilde{p} := \frac{\text{histogram}}{n}.$$

- Scale the received signal:

$$y := \frac{z}{\sqrt{En}} = \tilde{p} + \underbrace{\mathcal{N}\left(0, \frac{\sigma^2}{n^2 E}\right)}_{(*)}.$$

Questions:

- How bad is $(*)$?
- What estimator should be used?

Answers:

(i) $\tilde{p} \doteq \mathcal{N}(p_\theta, \frac{1}{n}\Sigma)$ for large n , where $\Sigma = \text{Diag}(p_\theta) - p_\theta p_\theta^T$.

$$y = \tilde{p} + \mathcal{N}(0, \frac{\sigma^2}{n^2 E}) \Rightarrow y \doteq \mathcal{N}(p_\theta, \frac{1}{n}\Sigma).$$

(ii) Maximum-likelihood estimator (MLE) based on y is prohibitive.

Let $y = \mathcal{N}(p_\theta, \frac{1}{n}\Sigma)$, then its pdf is

$$f(y_1, \dots, y_k \mid \theta) = \exp\left(-\frac{n \sum_{i=1}^k \frac{(p_\theta(i) - y_i)^2}{p_\theta(i)} + \log \prod_{i=1}^k p_\theta(i)}{2}\right) g(y).$$

Given y , minimize $\sum_{i=1}^k \frac{(p_\theta(i) - y_i)^2}{p_\theta(i)}$ for **asymptotic MLE**.

Theorem 1: The proposed estimator $\hat{\theta}$ minimizing

$$M(\theta) := \sum_{i=1}^k \frac{(p_{\theta}(i) - y_i)^2}{p_{\theta}(i)} \quad (3)$$

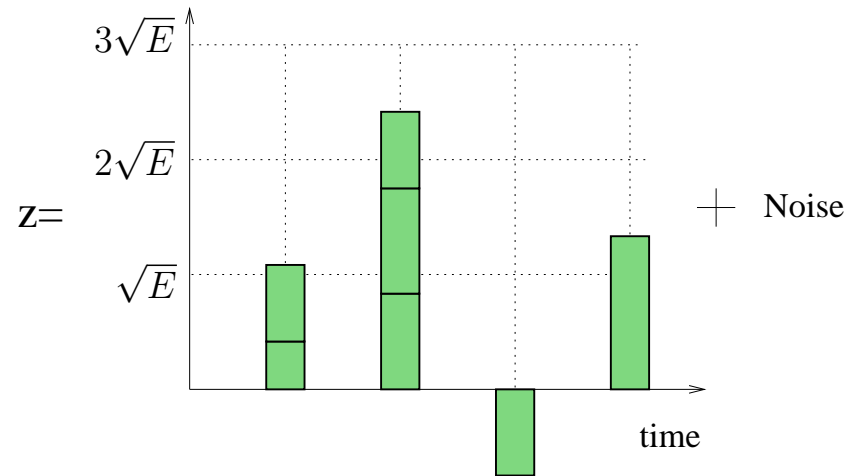
with respect to $\theta \in \mathbb{R}$ satisfies $\hat{\theta} \doteq \mathcal{N}(\theta, \frac{1}{nI(\theta)})$ for large n .

Remarks:

- The asymptotic performance of TBMA is as if the fusion center has direct access to X_i 's.
- No unbiased $\hat{\theta}$, even the ones with direct access to X_i 's, can do better than this.
- The theorem holds independent of the noise power σ^2 .
- The σ^2 determines the speed of convergence to the asymptotic MSE.

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TBMA with Random h_i



- Assume h_i has *non-zero* mean $h := \mathbb{E}(h_i)$, and $\sigma_h^2 := \text{Var}(h_i)$.
- Define $y = \frac{z}{\sqrt{E}hn}$.
- Observe $y \doteq \mathcal{N}(p_\theta, \frac{1}{n}\Sigma)$, where

$$\Sigma = \left(1 + \frac{\sigma_h^2}{h^2}\right) \text{Diag}(p_\theta) - p_\theta p_\theta^T.$$

$$f(y|\theta) \propto \exp\left(-\frac{n(y - p_\theta)^T \Sigma^{-1} (y - p_\theta) + \log |\Sigma|}{2}\right).$$

Theorem 2: The estimator $\hat{\theta}$ minimizing

$$M(\theta) = (y - p_\theta)^T \Sigma^{-1} (y - p_\theta), \quad (4)$$

with respect to $\theta \in \mathbb{R}$ satisfies

$$\hat{\theta} \doteq \mathcal{N}\left(\theta, \frac{1 + \frac{\sigma_h^2}{h^2}}{nI(\theta)}\right), \quad \text{for large } n.$$

Remarks:

- The performance loss due to channel randomness is $\propto (1 + \frac{\sigma_h^2}{h^2})$.
- When $h \approx 0$, the loss is significant.
- At the extreme case $h = 0$, the MSE does not go to zero even though $n \rightarrow \infty$. This is true for all unbiased $\hat{\theta}$.

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Performance of Orthogonal Allocation

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- Let $s_1, \dots, s_k \in \mathbb{C}^m$, $\|s_i\|^2 \leq 1$, be constellation points.
 - Sensor i transmits the waveform corresponding to $\sqrt{E}s_{X_i}$.
 - Signal received from the i 'th sensor:

$$z^{(i)} = h_i \sqrt{E} s_{X_i} + v^{(i)}, \quad (5)$$

where $v^{(i)} \sim \mathcal{N}(0, \sigma^2 I)$.

- When $h_i = 1$, the $z^{(1)}, \dots, z^{(n)}$ are i.i.d. with *Gaussian mixture* density:

$$z^{(i)} \sim \sum_{j=1}^k p_j \mathcal{N}(\sqrt{E} s_j, \sigma^2 I).$$

Asymptotic Performance:

- For any unbiased $\hat{\theta}$,

$$\mathbb{E}\{(\hat{\theta} - \theta)^2\} \geq \frac{1}{nJ(\theta)}, \quad (6)$$

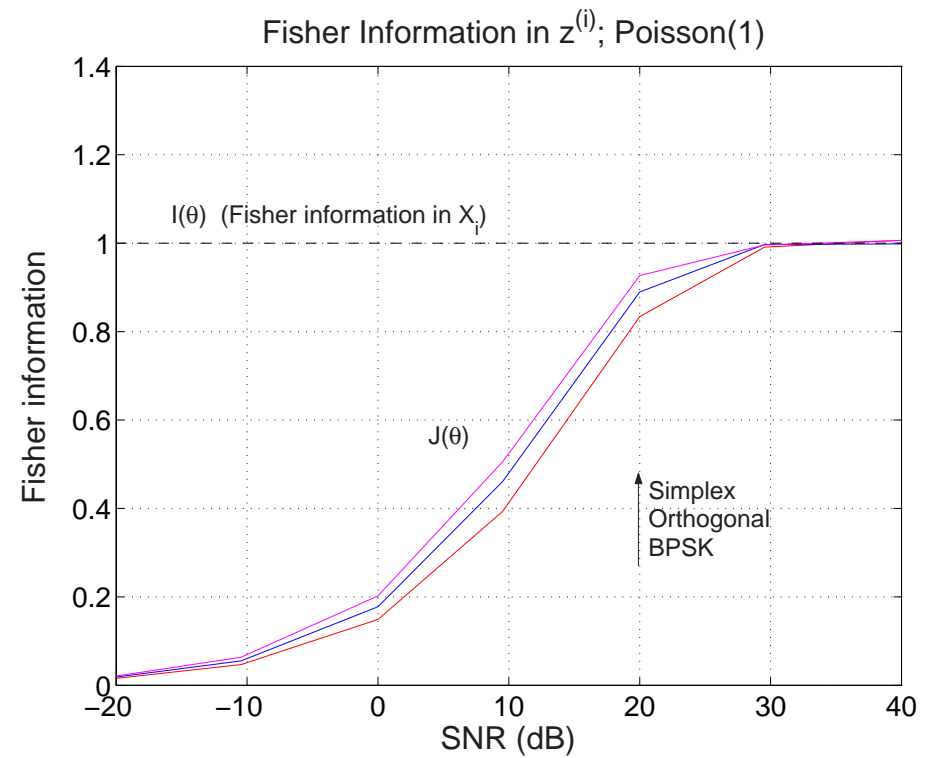
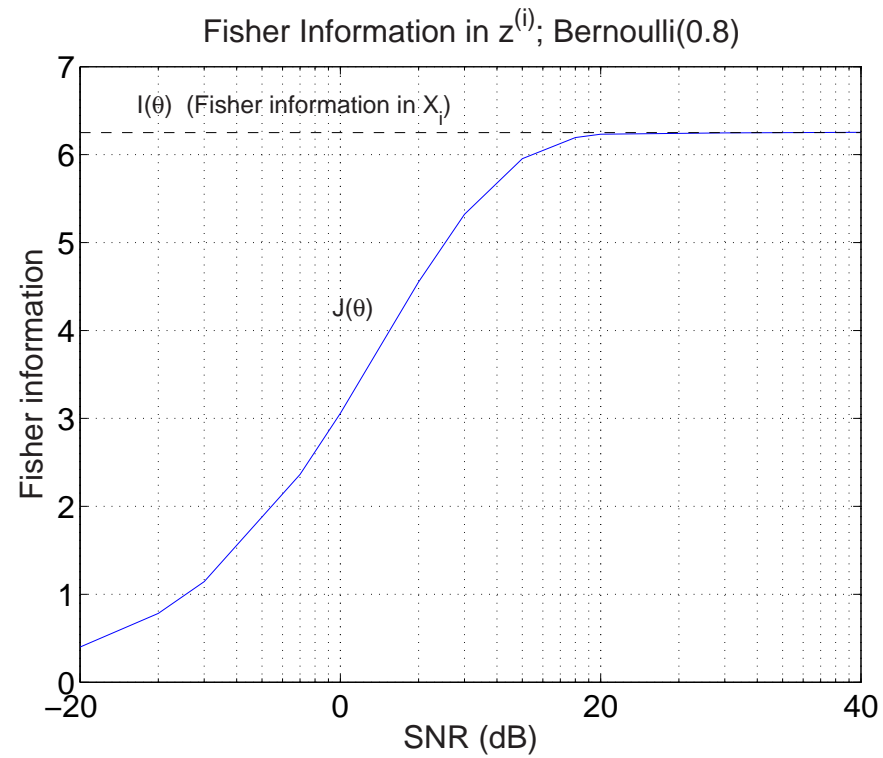
where $J(\theta) = \mathbb{E}_{z^{(i)}} \left[\left(\frac{d \log f(z^{(i)})}{d\theta} \right)^2 \right]$ is the *Fisher information* in $z^{(i)}$.

- The MLE based on $z^{(1)}, \dots, z^{(n)}$ satisfies $\hat{\theta} \doteq \mathcal{N}(\theta, \frac{1}{nJ(\theta)})$.

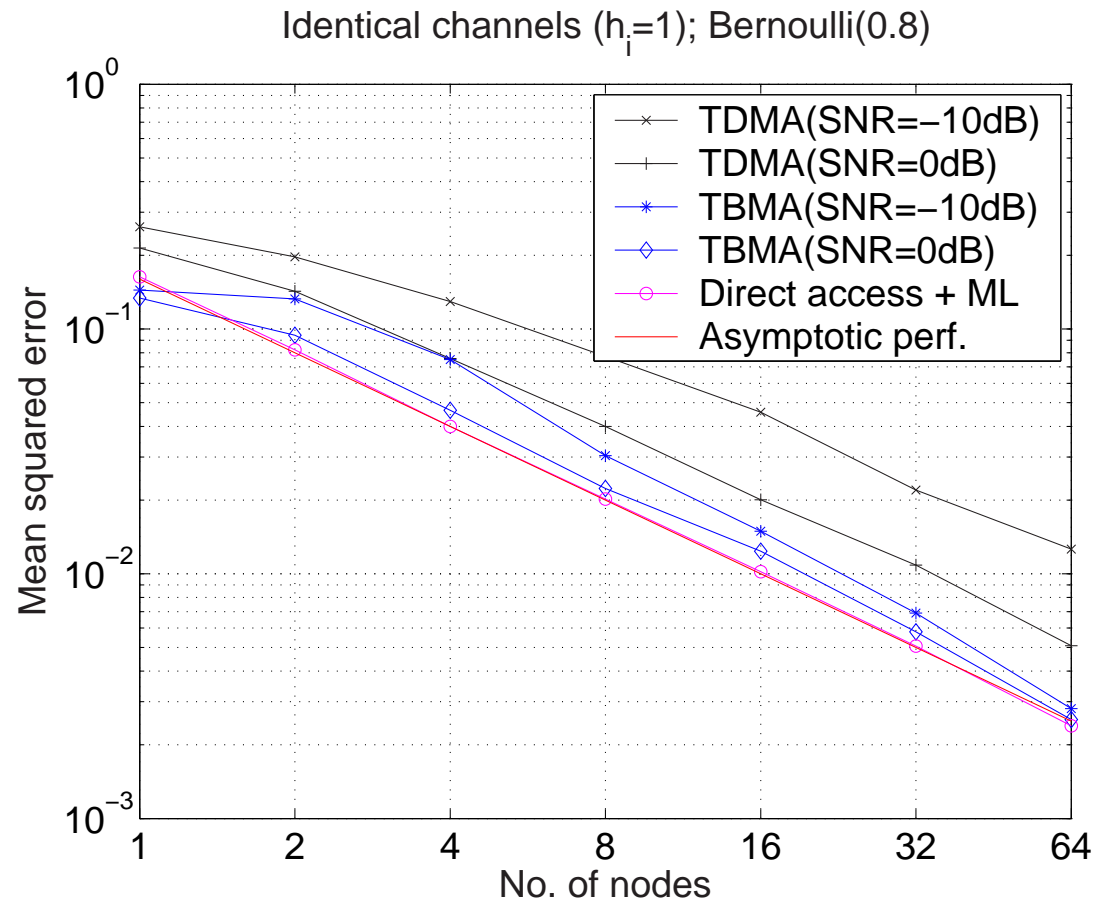
Remarks:

- For the best asymptotic performance $J(\theta)$ should be maximized with respect to s_1, \dots, s_k .
- For $h_i = 1$, the antipodal constellation maximizes $J(\theta)$ for $k = 2$.
- In general, the optimal constellation depends on the family $\{p_\theta : \theta \in \mathbb{R}\}$.

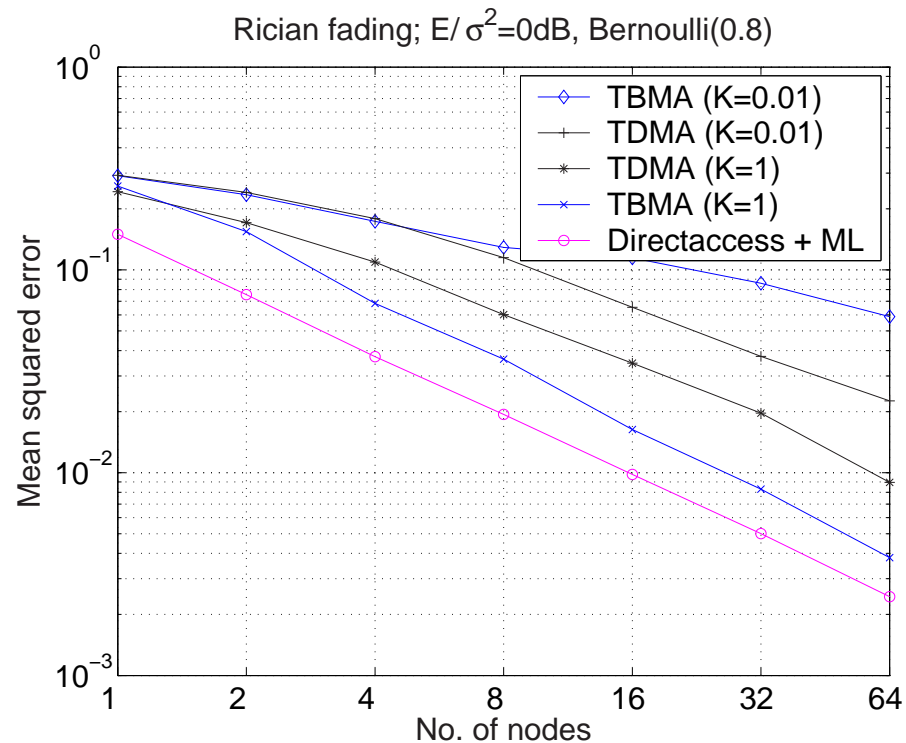
Performance of Orthogonal Allocation Cont'd



A Numerical Example



Numerical Example - 2



- The channel is Rician distributed, *i.e.*,

$$h_i = \sqrt{\frac{K}{K+1}} + \sqrt{\frac{1}{K+1}} \mathcal{CN}(0, 1),$$

where $K > 0$ is a deterministic number ($K = 0 \Rightarrow$ Rayleigh).

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Channel Side Information (CSI) at the Transmitter

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- In certain cases, the transmitter nodes may be able to learn their channel states before the transmission.
- Transmitter CSI can be utilized to solve the problem of zero-mean h_i .
- Let $h_i := r_i e^{j\rho_i}$.
- The i 'th node transmits $P(r_i) e^{-j\rho_i} \sqrt{E} \delta_{X_i}$ in TBMA, where $P(\cdot)$ is a **power control rule** satisfying

$$\mathbb{E}_{r_i}[P^2(r_i)] \leq 1.$$

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- One possibility is to invert the channel: $P(r) = \frac{1}{r}$. This effectively cancels out the effect of fading.
 - Complete inversion may require infinite energy $\mathbb{E}_r(1/r^2)$ as in the case of Rayleigh channels.
 - To circumvent this, consider the following generalization:

$$P(r) = \begin{cases} \alpha/r, & r \in [\beta, \infty) \\ \gamma, & \text{oth.} \end{cases}$$

where $\alpha, \beta, \gamma \in \mathbb{R}$ are constants independent of r .

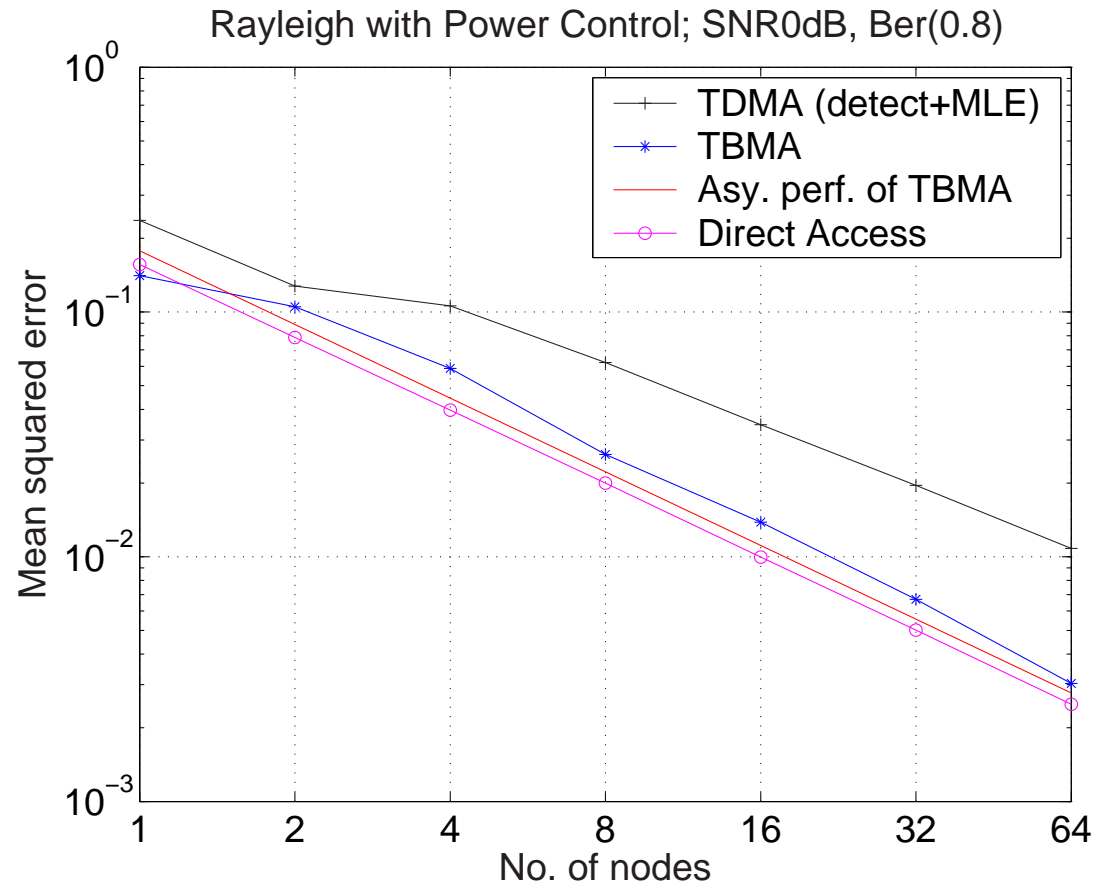
CSI at the Transmitter Cont'd

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- The channel gain seen by the receiver is $\tilde{h}_i := r_i P(r_i)$.
 - The asymptotic MSE is given by $(1 + \frac{\sigma_{\tilde{h}}^2}{\tilde{h}^2})/nI(\theta_0)$, where $\tilde{h} = \mathbb{E}(\tilde{h}_i)$, $\sigma_{\tilde{h}}^2 = \text{Var}(\tilde{h}_i)$.

Lemma: As α and β converge to zero,

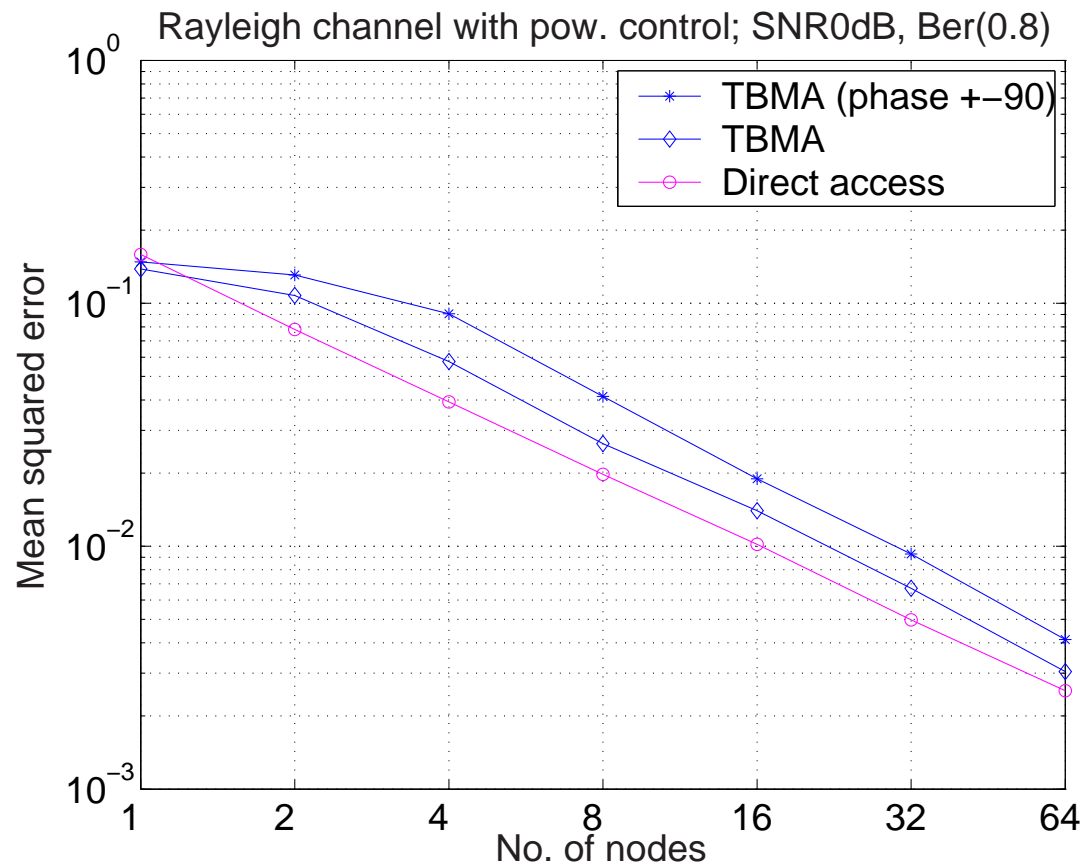
$$1 + \frac{\sigma_{\tilde{h}}^2}{\tilde{h}^2} \rightarrow 1.$$

Numerical Example - 3

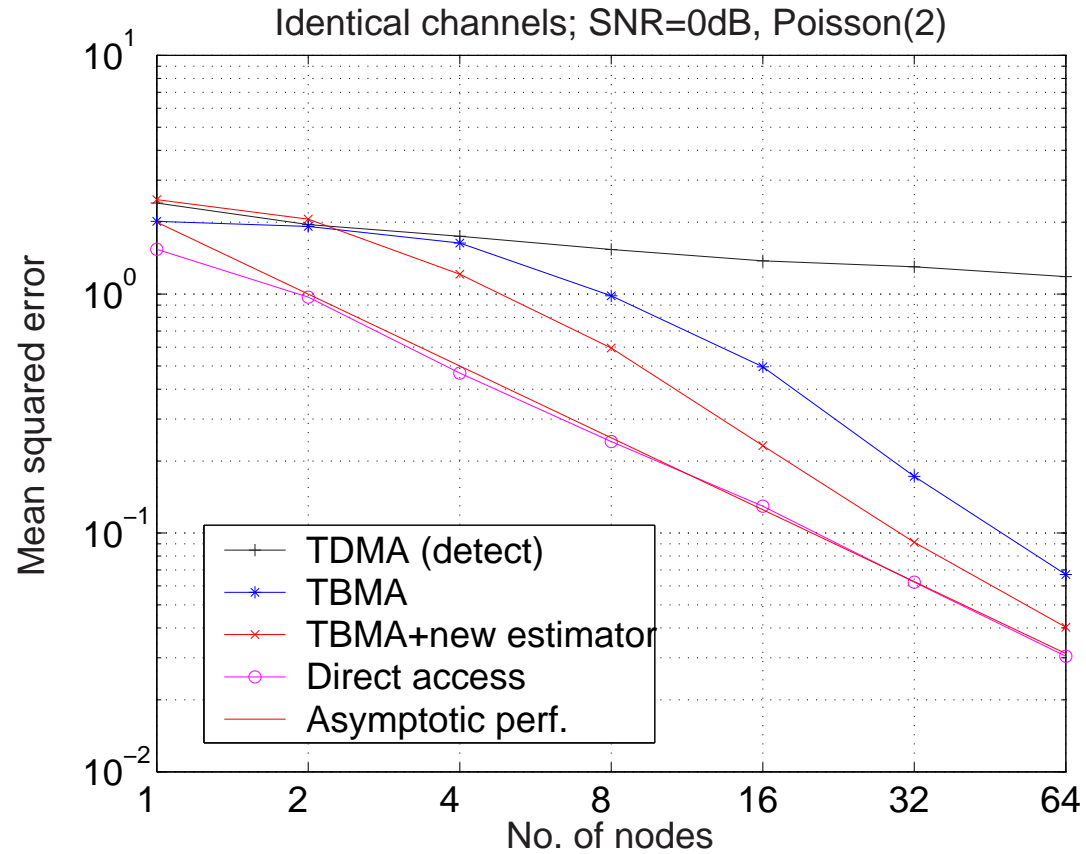


- Power controlled Rayleigh channel ($\alpha = 1, \beta = 0.89, \gamma = 1/\beta$).

Numerical Example - 4



Numerical Example - 5



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- We considered the problem of communicating sensor readings over a Gaussian multiaccess channel.
 - The TBMA scheme is proposed as an alternative to the conventional orthogonal allocation.
 - We proposed an estimator and characterized its asymptotic MSE.
 - The TBMA is bandwidth efficient and also has favorable MSE.
 - The TBMA needs to be used with transmitter CSI in case the channel has zero mean.

Related Publications

- G. Mergen and L. Tong "Type Based Estimation Over Multiaccess Channels," submitted to *IEEE Transactions on Signal Processing*, July 2004.
- G. Mergen and L. Tong "Estimation Over Deterministic Multiaccess Channels ," *42nd Annual Allerton Conference on Communications, Control and Computing*, 2004.
- K. Liu and A. Sayeed, "Optimal Distributed Detection Strategies for Wireless Sensor Networks," *42nd Annual Allerton Conference on Communications, Control and Computing*, Monticello, IL, 2004.